## **Text S1: the Wald test**

The full logistic model can be written as

$$\begin{split} \text{logit}(\pi_{gh}) &= \mu + \alpha_1 I(g = 1) + \alpha_2 I(g = 2) + \beta_1 I(h = 1) + \beta_2 I(h = 2) + \\ \lambda_{11} I(g = 1) I(h = 1) + \lambda_{12} I(g = 1) I(h = 2) + \lambda_{21} I(g = 2) I(h = 1) + \lambda_{22} I(g = 2) I(h = 2), \\ \text{where } \pi_{gh} \text{ is the probability of disease given genotype } g \text{ at SNP1 and } h \text{ at SNP2.} \end{split}$$

The log-likelihood is  $l = \sum_{g=0}^{2} \sum_{h=0}^{2} \left[ n_{gh1} \log(\pi_{gh}) + n_{gh0} \log(1 - \pi_{gh}) \right]$ . It is easy to verify that the

MLEs of  $\pi_{gh}$  are  $n_{gh1} / n_{gh+}$ , where  $n_{gh+} = n_{gh0} + n_{gh1}$ . Therefore, the MLEs of parameters

 $\mu, \alpha_{g}, \beta_{h}, \lambda_{gh}$  are respectively

$$\begin{aligned} \hat{\mu} &= \operatorname{logit}(\hat{\pi}_{00}) = \operatorname{log}(n_{001} / n_{000}) \\ \hat{\alpha}_{g} &= \operatorname{logit}(\hat{\pi}_{g0}) - \operatorname{logit}(\hat{\pi}_{00}) = \operatorname{log}\left(\frac{n_{g01}}{n_{001}} / \frac{n_{g00}}{n_{000}}\right), g = 1,2 \\ \hat{\beta}_{h} &= \operatorname{logit}(\hat{\pi}_{0h}) - \operatorname{logit}(\hat{\pi}_{00}) = \operatorname{log}\left(\frac{n_{0h1}}{n_{001}} / \frac{n_{0h0}}{n_{000}}\right), h = 1,2 \\ \hat{\lambda}_{gh} &= \operatorname{logit}(\hat{\pi}_{gh}) + \operatorname{logit}(\hat{\pi}_{00}) - \operatorname{logit}(\hat{\pi}_{g0}) - \operatorname{logit}(\hat{\pi}_{0h}) \\ &= \operatorname{log}\left(\frac{n_{gh1}n_{001}}{n_{g01}n_{0h1}} / \frac{n_{gh0}n_{000}}{n_{g00}n_{0h0}}\right), g = 1,2, h = 1,2 \end{aligned}$$

Let 
$$a_{gh} = \frac{n_{gh0}n_{gh1}}{n_{gh0} + n_{gh1}}, a_{g+} = \sum_{h=0}^{2} a_{gh}, a_{+h} = \sum_{g=0}^{2} a_{gh}, a_{++} = \sum_{g=0}^{2} \sum_{h=0}^{2} a_{gh}$$
, where

g = 0,1,2; h = 0,1,2.

The observed Fisher Information matrix is

$$I = \begin{pmatrix} a_{++} & a_{1+} & a_{2+} & a_{+1} & a_{+2} & a_{11} & a_{12} & a_{21} & a_{22} \\ a_{1+} & a_{1+} & 0 & a_{11} & a_{12} & a_{11} & a_{12} & 0 & 0 \\ a_{2+} & 0 & a_{2+} & a_{21} & a_{22} & 0 & 0 & a_{21} & a_{22} \\ a_{+1} & a_{11} & a_{21} & a_{+1} & 0 & a_{11} & 0 & a_{21} & 0 \\ a_{+2} & a_{12} & a_{22} & 0 & a_{+2} & 0 & a_{12} & 0 & a_{22} \\ a_{11} & a_{11} & 0 & a_{11} & 0 & a_{11} & 0 & 0 & 0 \\ a_{12} & a_{12} & 0 & 0 & a_{12} & 0 & a_{12} & 0 & 0 \\ a_{21} & 0 & a_{21} & a_{21} & 0 & 0 & 0 & a_{21} & 0 \\ a_{22} & 0 & a_{22} & 0 & a_{22} & 0 & 0 & 0 & a_{22} \end{pmatrix}$$

By the asymptotic properties of MLEs, the distribution of the MLE vector  $\hat{\theta}$  is approximately multivariate normal with variance-covariance matrix  $(I)^{-1}$ . Therefore, the joint distribution of the MLEs for the interaction parameters  $\hat{\theta}_{\lambda} = (\hat{\lambda}_{11}, \hat{\lambda}_{12}, \hat{\lambda}_{21}, \hat{\lambda}_{22})^T$  is approximately multivariate normal with variance-covariance matrix being the elements in the last four rows and columns of  $(I)^{-1}$ . Using symbolic matrix inverse in Mathematica, we found that the covariance-covariance matrix is

$$\hat{cov}(\hat{\theta}_{\lambda}) = \begin{pmatrix} \frac{1}{a_{00}} + \frac{1}{a_{10}} + \frac{1}{a_{10}} + \frac{1}{a_{11}} & \frac{1}{a_{00}} + \frac{1}{a_{10}} & \frac{1}{a_{10}} + \frac{1}{a_{10}} & \frac{1}{a_{00}} + \frac{1}{a_{00}} & \frac{1}{a_{00}} \\ \frac{1}{a_{00}} + \frac{1}{a_{10}} & \frac{1}{a_{00}} + \frac{1}{a_{10}} + \frac{1}{a_{10}} + \frac{1}{a_{12}} & \frac{1}{a_{00}} & \frac{1}{a_{00}} + \frac{1}{a_{00}} + \frac{1}{a_{02}} \\ \frac{1}{a_{00}} + \frac{1}{a_{01}} & \frac{1}{a_{00}} + \frac{1}{a_{00}} & \frac{1}{a_{00}} + \frac{1}{a_{00}} + \frac{1}{a_{00}} + \frac{1}{a_{20}} + \frac{1}{a_{21}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} \\ \frac{1}{a_{00}} & \frac{1}{a_{00}} + \frac{1}{a_{02}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} \\ \frac{1}{a_{00}} & \frac{1}{a_{00}} + \frac{1}{a_{02}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} \\ \frac{1}{a_{00}} & \frac{1}{a_{00}} + \frac{1}{a_{02}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} \\ \frac{1}{a_{00}} & \frac{1}{a_{00}} + \frac{1}{a_{02}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} \\ \frac{1}{a_{00}} & \frac{1}{a_{00}} + \frac{1}{a_{00}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} \\ \frac{1}{a_{00}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} & \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} \\ \frac{1}{a_{00}} & \frac{1}{a_{00}} + \frac{1}{a_{20}} & \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} \\ \frac{1}{a_{20}} & \frac{1}{a_{20}} + \frac{1}{a_{20}} + \frac{1}{a_{20}} & \frac{1}{a_{20}} + \frac{1}{a_{20}} \\ \frac{1}{a_{20}} & \frac{1}{a_{20}} + \frac{1}{a_{20}}$$

Note that the covariance matrix can also be derived using the multivariate normal approximation to the distribution of the observed counts.

The Wald test for testing  $H_0: \lambda_{11} = \lambda_{12} = \lambda_{21} = \lambda_{22} = 1$  is given by

$$Wald = \hat{\vec{\theta}}_{\lambda}^{T} \left[ \hat{cov}(\hat{\vec{\theta}}_{\lambda}) \right]^{-1} \hat{\vec{\theta}}_{\lambda}$$

Under the null hypothesis of no gene-gene interaction, *Wald* follows the chi-squared distribution with 4 *df* asymptotically.