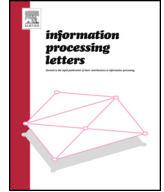




Contents lists available at ScienceDirect

Information Processing Letters

journal homepage: www.elsevier.com/locate/ipl

Simplified Chernoff bounds with powers-of-two probabilities

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ARTICLE INFO

Article history:

Received 20 September 2022

Received in revised form 17 February 2023

Accepted 2 April 2023

Available online 5 April 2023

Keywords:

Algorithms

Analysis of algorithms

Algorithm analysis

Chernoff bound

Lambert W function

ABSTRACT

In this paper, we derive simplified Chernoff bounds with powers-of-two probabilities, and we show their uses in analyzing probabilistic algorithms.

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1. Introduction

Chernoff bounds [4,11] have been shown to be useful for analyzing a wide variety of different probabilistic algorithms and processes, e.g., see [1,9,12,13].

Suppose X_1, X_2, \dots, X_n are independent random variables taking values in $\{0, 1\}$. Let $X = \sum_{i=1}^n X_i$ and let $\mu = E[X]$ denote X 's expected value. In their more general (multiplicative) forms for such a random variable, X , Chernoff bounds can be stated as follows, e.g., see [1,9,12–14].

Theorem 1. For any $\delta > 0$,

$$\Pr(X > (1 + \delta)\mu) < \left(\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right)^\mu.$$

Also, for any $0 < \delta < 1$,

$$\Pr(X < (1 - \delta)\mu) < \left(\frac{e^{-\delta}}{(1 - \delta)^{1-\delta}} \right)^\mu.$$

These formulas are unwieldy to use in practice, however; hence, researchers often use other forms of the Chernoff bounds, with the following being common (see, e.g., [1,2,9,12–14]):

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Theorem 2.

$$\Pr(X > (1 + \delta)\mu) < e^{-\delta^2\mu/(2+\delta)}, \quad \text{for } \delta > 0, \quad (1)$$

$$\Pr(X < (1 - \delta)\mu) < e^{-\delta^2\mu/2}. \quad \text{for } 0 < \delta < 1. \quad (2)$$

As evidence for how influential these bounds have been, we note that a 1990 paper from *Information Processing Letters*, by Hagerup and Rüb [9], which includes bounds like those in Theorem 2, has been cited over 700 times! Indeed, these bounds have become so well-known that researchers often use them without citation.

As in Theorem 2, probabilities in simplified Chernoff bounds are typically expressed as powers of Euler's number, e , whereas in Computer Science applications it is often preferred to express probabilities in terms of powers of 2, for which simplified Chernoff bounds are lacking. Indeed, some researchers apply a Chernoff bound, as in Theorem 2, and then convert the resulting probability to a power of two using the crude inequality, $2 \leq e$. For example, see Elsässer and Sauerwald [8]. Of course, one can use a slightly better inequality to derive the following.

Corollary 3.

$$\Pr(X > (1 + \delta)\mu) < 2^{-1.442\delta^2\mu/(2+\delta)} < 2^{-7\delta^2\mu/(10+5\delta)}, \quad \text{for } \delta > 0, \quad (3)$$

$$\Pr(X < (1 - \delta)\mu) < 2^{-1.442\delta^2\mu/2} < 2^{-7\delta^2\mu/10}, \quad \text{for } 0 < \delta < 1. \quad (4)$$

Proof. Note that $2^{7/5} < e$, since $\log_2 e \approx 1.442695$; hence, the bounds follow immediately from Theorem 2. ■

In this paper, we are interested in simplified Chernoff bounds with powers-of-two probabilities for reasonable values of δ , as such bounds are often used in Computer Science applications. In terms of prior work, there is a notable upper-tail power-of-two Chernoff bound from a book by Mitzenmacher and Upfal [12] and the *IPL* paper by Hagerup and Rüb [9]:

Theorem 4 ([12] (p. 69) and [9]).

$$\Pr(X > R) < 2^{-R}, \quad \text{for } R \geq 6\mu.$$

In addition, Motwani and Raghavan [13] leave as an exercise to prove $\Pr(X > R) < 2^{-R}$, for $R \geq 2e\mu$, which is a slightly better condition, since $2e \approx 5.43656$. Although this and the power-of-two Chernoff bound of Theorem 4 are useful, we show below that $\Pr(X > R) < 2^{-R}$, for $R \geq 4.5\mu$, which can lead to better analyses for randomized algorithms. Indeed, in this paper, we derive a number of such simplified Chernoff bounds with powers-of-two probabilities, for both upper and lower tails, for reasonable values of δ . We also mention some applications of our simplified powers-of-two Chernoff bounds, but these are just the tip of the iceberg in terms of improved analyses of algorithms that are possible, e.g., given that the *IPL* paper by Hagerup and Rüb [9] has been cited over 700 times.

2. The Lambert W function

Some of our proofs make use of the Lambert W function; hence, before we derive our simplified powers-of-two Chernoff bounds, let us first review this function. The Lambert W function is defined by the rule that $W(z) = w$ iff w satisfies the equation

$$we^w = z,$$

e.g., see Corless, Gonnet, Hare, Jeffrey, and Knuth [5] or Corless, Jeffrey, and Knuth [6]. Technically, W is not a function. Hence its real-valued expression is partitioned into two branches: $W_0(x)$, which is called the *principal branch* and is always greater than or equal to -1 , and $W_{-1}(x)$, which is called the *non-principal branch* and is always less than or equal to -1 . A plot of the two real branches is shown in Fig. 1. The two branches split at $(-\frac{1}{e}, -1)$. $W_0(ye^y) = y$ for $y \geq -1$, and $W_{-1}(ye^y) = y$ for $y \leq -1$.

The Lambert W function has many applications, including characterizing the number of unrooted trees [5]. It cannot be expressed in terms of elementary functions; hence, evaluating it typically requires one to use a numerical algorithm, e.g., see [3].

3. Improved powers-of-two Chernoff bounds

In this section, we derive simplified Chernoff bounds with powers-of-two probabilities.

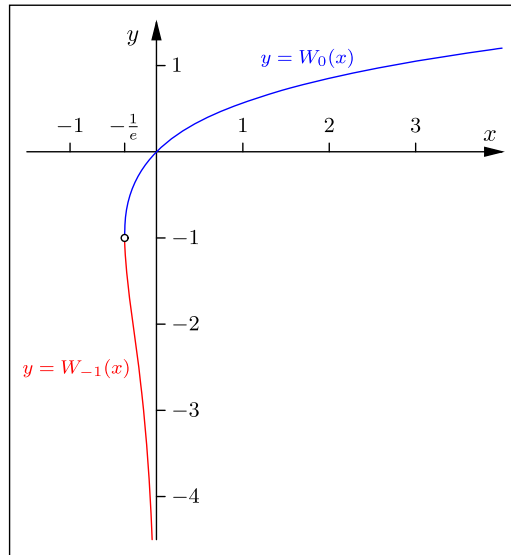


Fig. 1. The two real branches of the Lambert W function. Image Copyright © 2022 Michael Dillencourt; used with permission.

3.1. Upper-tail bounds

We begin with some upper-tail bounds. The first is a strict improvement of Theorem 4.

Theorem 5.

$$\Pr(X > R) < 2^{-R}, \quad \text{for } R \geq 4.5\mu.$$

Proof. We actually show that $\Pr(X > R) < 2^{-R}$, for $R \geq 4.31107\mu > -\mu/W_0(-1/2e)$, where $W_0(x)$ is the principal branch of the Lambert W function. From the general form of the Chernoff bound of Theorem 1, taking $R = (1 + \delta)\mu$,

$$\Pr(X > R) = \Pr(X > (1 + \delta)\mu) < \left(\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right)^\mu.$$

In order for this probability to be at most 2^{-R} , we need $1 + \delta \geq 2e^{\delta/(1+\delta)}$. Setting $x = 1 + \delta$, the breakpoint for this inequality occurs for x satisfying

$$x = 2e^{\frac{x-1}{x}}.$$

That is,

$$xe^{-\frac{x-1}{x}} = 2.$$

Putting this into the form of the Lambert W function definition, let $u = -(x - 1)/x$, so $x = 1/(1 + u)$. The equation becomes

$$\left(\frac{1}{1 + u} \right) e^u = 2,$$

which can be rewritten as

$$-(1 + u)e^{-(1+u)} = -\frac{1}{2e}.$$

This equation has $u = -W_0(-1/2e) - 1$ as a solution. After back-substituting the solution is $x = -1/W_0(-1/2e)$. Numerically, $-1/W_0(-1/2e) \approx -1/(-0.23196) \approx 4.31107$, which establishes the bound for R . ■

We can also establish the following general bounds.

Theorem 6. The bound

$$\Pr(X > (1 + \delta)\mu) < 2^{-\alpha\mu}, \tag{5}$$

holds:

1. For fixed $\delta > 0$ when

$$\alpha \leq \log_2 \left(\frac{(1 + \delta)^{1+\delta}}{e^\delta} \right). \tag{6}$$

2. For fixed $\alpha > 0$ when

$$\delta \geq e^{W_0\left(\frac{\alpha \ln 2 - 1}{e}\right) + 1} - 1. \tag{7}$$

Proof. By Theorem 1, (5) holds whenever we have:

$$2^{-\alpha\mu} \geq \left(\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right)^\mu. \tag{8}$$

Part 1 follows from the observation that if (6) holds, so does (8). To establish part 2, we fix δ and determine the conditions for which (8) holds. (8) holds when

$$\left(\frac{1 + \delta}{e} \right)^{1+\delta} \geq \frac{2^\alpha}{e}.$$

Since $\ln x$ and x/e are both monotone increasing functions of x for positive x , this is equivalent to

$$\left(\frac{1 + \delta}{e} \right) \ln \left(\frac{1 + \delta}{e} \right) \geq \frac{\alpha \ln 2 - 1}{e}.$$

Since the left-hand side is of the form xe^x where $x = \ln((1 + \delta)/e)$, and since $\ln((1 + \delta)/e) > -1$ for any positive δ , we can rewrite the last equation as

$$\ln \left(\frac{1 + \delta}{e} \right) \geq W_0 \left(\frac{\alpha \ln 2 - 1}{e} \right),$$

from which (7) follows. ■

We also have the following theorem, which provides a bound for modest values of δ .

Theorem 7.

$$\Pr(X > (1 + \delta)\mu) < 2^{-\delta\mu}, \quad \text{for } \delta > 2.20603. \tag{9}$$

Proof. From the general form of the Chernoff bound of Theorem 1,

$$\Pr(X > (1 + \delta)\mu) < \left(\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right)^\mu.$$

For this probability to be at most $2^{-\delta\mu}$, we need $2^{-\delta}(1 + \delta)^{1+\delta} \geq e^\delta$. This can be rewritten as

$$\left(\frac{1 + \delta}{2e} \right)^{1+\delta} \geq \frac{1}{2e}. \tag{10}$$

Taking both sides to the $1/(2e)$ power, and taking the log of both sides yields

$$\left(\frac{1 + \delta}{2e} \right) \ln \left(\frac{1 + \delta}{2e} \right) \geq \left(\frac{1}{2e} \right) \ln \left(\frac{1}{2e} \right). \tag{11}$$

We can see by inspection that one breakpoint for (11) is $\delta = 0$, which is the trivial value for which equality holds in (10). Since $\ln(1/(2e)) < -1$, this corresponds to choosing the non-principal branch of the Lambert function. Choosing the principal branch, we see that equality occurs in (11) when

$$W_0 \left(\frac{-\ln 2 - 1}{2e} \right) = \ln \left(\frac{1 + \delta}{2e} \right).$$

This produces the solution

$$\delta = 2e \cdot e^{W_0\left(\frac{-\ln 2 - 1}{2e}\right) - 1} \approx 2e \cdot e^{W_0(0.31144) - 1} \approx 2e \cdot e^{-0.52811} - 1 \approx 2.20603. \quad \blacksquare$$

We can use Theorem 6 to derive the following specific upper-tail powers-of-two Chernoff bounds for smaller values of δ , all of which are tighter than the bounds of Corollary 3.

Corollary 8.

$$\Pr(X > (1 + \delta)\mu) < 2^{-\mu}, \quad \text{for } \delta \geq 1.4, \tag{12}$$

$$\Pr(X > (1 + \delta)\mu) < 2^{-0.557\mu} < 2^{-5\mu/9}, \quad \text{for } \delta \geq 1, \tag{13}$$

$$\Pr(X > (1 + \delta)\mu) < 2^{-0.266\mu} < 2^{-\mu/4}, \quad \text{for } \delta \geq 2/3, \tag{14}$$

$$\Pr(X > (1 + \delta)\mu) < 2^{-0.156\mu} < 2^{-3\mu/20}, \quad \text{for } \delta \geq 1/2, \tag{15}$$

$$\Pr(X > (1 + \delta)\mu) < 2^{-0.072\mu} < 2^{-\mu/14}, \quad \text{for } \delta \geq 1/3, \tag{16}$$

$$\Pr(X > (1 + \delta)\mu) < 2^{-0.0417\mu} < 2^{-\mu/24}, \quad \text{for } \delta \geq 1/4. \tag{17}$$

Proof. To derive (12), we set $\alpha = 1$ and use part 2 of Theorem 6. By (7), the minimum value of δ is

$$e^{W_0\left(\frac{\ln 2 - 1}{e}\right) + 1} - 1 \approx e^{W_0(-0.11288) + 1} - 1 \approx e^{-0.128337 + 1} - 1 \approx 1.39088 < 1.4.$$

To derive (13), we set $\delta = 1$ and use part 1 of Theorem 6. By (6), the maximum value of α is

$$\log_2\left(\frac{2^2}{e^2}\right) \approx 0.557.$$

The derivations of (14) through (17) are similar to that of (13). ■

3.2. Lower-tail bounds

We also derive lower-tail Chernoff bounds that improve the Chernoff bounds of Corollary 3. We first prove the following analog of Theorem 6.

Theorem 9. *The bound*

$$\Pr(X < (1 - \delta)\mu) < 2^{-\beta\mu}, \tag{18}$$

holds:

1. For fixed δ with $0 < \delta < 1$ when

$$\beta \leq \log_2(e^\delta(1 - \delta)^{1-\delta}). \tag{19}$$

2. For fixed $\beta > 0$ when

$$1 - e^{W_{-1}\left(\frac{\beta \ln 2 - 1}{e}\right) + 1} \leq \delta < 1. \tag{20}$$

Proof. By Theorem 1, (18) holds whenever we have:

$$2^{-\beta\mu} \geq \left(\frac{e^{-\delta}}{(1 - \delta)^{1-\delta}}\right)^\mu \tag{21}$$

Part 1 follows from the observation that if (19) holds, so does (21). To establish part 2, we fix β and determine the values of δ between 0 and 1 for which (21) holds. (21) holds when

$$\left(\frac{1 - \delta}{e}\right)^{1-\delta} \geq \frac{2^\beta}{e}. \tag{22}$$

Since the functions $\ln x$ and x/e are both monotone increasing functions of x for positive x and since both quantities in (22) are less than 1, this is equivalent to

$$\left(\frac{1 - \delta}{e}\right) \ln\left(\frac{1 - \delta}{e}\right) \leq \frac{\beta \ln 2 - 1}{e}.$$

Since the left-hand side is of the form xe^x where $x = \ln((1 - \delta)/e)$, and since $\ln((1 - \delta)/e) < -1$ for any positive δ , we can rewrite the last equation as

$$\ln\left(\frac{1-\delta}{e}\right) \leq W_{-1}\left(\frac{\beta \ln 2 - 1}{e}\right),$$

from which (20) follows. ■

Corollary 10.

$$\Pr(X < (1-\delta)\mu) < 2^{-\mu}, \quad \text{for } 0.9099 \leq \delta < 1. \tag{23}$$

Proof. To derive (23), we set $\beta = 1$ and use part 2 of Theorem 9. By (20), the minimum value of δ is

$$1 - e^{W_{-1}\left(\frac{\ln 2 - 1}{e}\right) + 1} \approx 1 - e^{W_{-1}(-0.11288) + 1} - 1 \approx 1 - e^{-3.40737 + 1} \approx 0.9099. \quad \blacksquare$$

If we are interested in bounds for slightly smaller values of δ , we can use the following theorem.

Theorem 11.

$$\Pr(X < (1-\delta)\mu) < 2^{-\delta\mu}, \quad \text{for } 0.8687 \leq \delta < 1. \tag{24}$$

Proof. From the general form of the Chernoff bound of Theorem 1,

$$\Pr(X < (1-\delta)\mu) < \left(\frac{e^{-\delta}}{(1-\delta)^{1-\delta}}\right)^\mu.$$

For this probability to be at most $2^{-\delta\mu}$, we need $e^\delta(1-\delta)^{1-\delta} \geq 2^\delta$. This can be rewritten as

$$\left(\frac{2(1-\delta)}{e}\right)^{1-\delta} \geq \frac{2}{e}. \tag{25}$$

Taking both sides to the $2/e$ power, and taking the log of both sides yields

$$\left(\frac{2(1-\delta)}{e}\right) \ln\left(\frac{2(1-\delta)}{e}\right) \geq \left(\frac{2}{e}\right) \ln\left(\frac{2}{e}\right). \tag{26}$$

We can see by inspection that one breakpoint for (26) is $\delta = 0$, which is the trivial value for which equality holds in (25). Since $\ln(2/e) > -1$, this corresponds to choosing the principal branch of the Lambert function. Choosing the non-principal branch, we see that equality occurs in (26) when

$$W_{-1}\left(\frac{2(\ln 2 - 1)}{e}\right) = \ln\left(\frac{2(1-\delta)}{e}\right).$$

This produces the solution

$$\delta = 1 - \frac{e^{W_{-1}\left(\frac{2(\ln 2 - 1)}{e}\right) + 1}}{2} \approx 1 - \frac{e^{W_{-1}(-0.2257) + 1}}{2} \approx 1 - \frac{e^{-2.3372 + 1}}{2} \approx 0.8687. \quad \blacksquare$$

We can also derive the following specific lower-tail powers-of-two Chernoff bounds for smaller values of δ , all of which are tighter than the bounds of Corollary 3.

Corollary 12. Suppose $\delta < 1$. Then

$$\Pr(X < (1-\delta)\mu) < 2^{-0.771\mu} < 2^{-3\mu/4}, \quad \text{for } \delta \geq 5/6, \tag{27}$$

$$\Pr(X < (1-\delta)\mu) < 2^{-0.582\mu} < 2^{-5\mu/9}, \quad \text{for } \delta \geq 3/4, \tag{28}$$

$$\Pr(X < (1-\delta)\mu) < 2^{-0.433\mu} < 2^{-3\mu/7}, \quad \text{for } \delta \geq 2/3, \tag{29}$$

$$\Pr(X < (1-\delta)\mu) < 2^{-0.221\mu} < 2^{-\mu/5}, \quad \text{for } \delta \geq 1/2, \tag{30}$$

$$\Pr(X < (1-\delta)\mu) < 2^{-0.09092\mu} < 2^{-\mu/11}, \quad \text{for } \delta \geq 1/3, \tag{31}$$

$$\Pr(X < (1-\delta)\mu) < 2^{-0.049\mu} < 2^{-\mu/21}, \quad \text{for } \delta \geq 1/4. \tag{32}$$

Proof. To derive (27), we set $\delta = 5/6$ and use part 1 of Theorem 9. By (19), the maximum value of β is

$$\log_2\left(e^{5/6}(1/6)^{1/6}\right) \approx 0.7714.$$

The derivations of (28) through (32) are similar to that of (27). ■

4. Applications

In this section, we highlight some improved analyses that are implied by the above simplified Chernoff bounds.

Hassin and Peleg [10] study a probabilistic local polling process, examine its properties, and propose its use in the context of distributed network protocols for achieving consensus. Their analysis uses Theorem 4 to show that a parallel random-walk process will succeed with half of the pairs of random walks meeting in $41M$ expected steps, where M is the maximum expected meeting time for two walks. Substituting Theorem 5 in their analysis improves the expected number of steps for half of the pairs meeting to $24M$.

Diks and Pelc [7] present an algorithm to exchange values between all fault-free nodes in an n -node network where nodes and links fail with constant probabilities, basing their analysis, in part, on Theorem 4. Substituting Theorem 5 in their analysis improves the constant factor in their analysis and/or the probability of failure that their algorithm tolerates.

Elsässer and Sauerwald [8] study a randomized broadcasting protocol. Their analysis uses Theorem 2, with $\delta = 5/6$, and the crude inequality $2 < e$ to bound the failure probability of their algorithm to be at most $1/n$. Simply substituting Theorem 12 in their analysis improves their failure probability to $n^{-3.5}$.

Declaration of competing interest

The authors declare the following financial interests/personal relationships which may be considered as potential competing interests: Michael Goodrich reports financial support was provided by National Science Foundation (2212129). Michael Goodrich reports a relationship with National Science Foundation (2212129) that includes: funding grants.

Data availability

No data was used for the research described in the article.

Acknowledgements

This research was supported in part by the National Science Foundation under grant 2212129.

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